

## Can ABC Lead to Sustained 123? The Medium-Term Effects of a Technology-Enhanced Adult Education Program

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**Abstract.** Can information technology preserve the short-term learning gains associated with adult education programs? This study estimates the medium-term impacts of a mobile phone module (Project ABC) that was added to a standard adult education curriculum, and for which there were significant short-term impacts on educational outcomes. Two years after the end of the program, students in ABC villages had reading scores that were significantly higher than those in standard adult education classes, and women and younger students were better able to decode numbers. This can be partially attributed to more active mobile phone usage in ABC villages. Households in ABC villages also were more likely to own certain durable assets, had lower levels of food security and were more likely to save. Overall, these results suggest that short-term learning gains associated with technology can persist, especially if students have the opportunity to practice with that technology after the end of classes.

Keywords: Education; Information Technology; Niger

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## **I. Introduction**

The question of whether educational gains persist is of interest to academics and policymakers alike. The belief that educational investments will lead to sustained improvements in learning, and hence in improvements in adults' living standards, underlies many education programs. Yet despite improvements in primary school enrolment since the 1990s, there have been limited improvements in learning outcomes. This, as well as the widespread growth of information technology in many developing countries, has been the rationale for using "technology-assisted" instruction, in the hopes that such technology would lead to improvements in learning in the short- and long-term.

Existing research on the impact of technology-assisted learning has mixed results, with positive impacts of computer-assisted instruction on learning in India and China (Banerjee et al 2007, Linden 2008, Lai et al 2011, 2013, Mo et al 2014, Deshpande et al 2017) to no effect in Colombia, Peru and the U.S. (Barrera-Osorio and Linden 2009, Beuermann et al. 2013, Fairlie and Robinson 2014). Focusing on mobile phone technology and adult learning, Aker et al (2012) find that a mobile phone-enhanced adult education program in Niger significantly improved adults' reading and math z-scores in the short-term. Similarly, Aker et al (2017) find that a mobile phone-based adult education intervention in Los Angeles significantly improved adults' reading levels in Spanish over a four-month period.

While there is significant research showing that learning gains from primary education fall rapidly (ie, Kane and Staiger 2008, Jacob, Lefgren, and Sims 2010, Rothstein 2010, Glewwe et al 2010, Banerjee et al. 2007), there are few studies in the area of technology-assisted learning. One notable exception is that of Banerjee et al

(2007), who find that learning gains associated with a computer-assisted education program fell rapidly within one year of the program (Banerjee et al 2007). Yet, to our knowledge, there is little evidence in economics on the impact of technology-assisted instruction on adults' learning outcomes in the medium-term, nor on the impact of such programs on other indicators of well-being.

In prior work, Aker et al (2012) measured the impact of a mobile phone-assisted module (ABC) as part of an adult education program, finding that simply learning how to use a mobile phone increased adults reading and math z-scores by .19-.25 s.d., with persistent effects seven months after the program. The impacts were similar by gender and age, although relatively stronger in one region, and were primarily driven by students' increased use of the mobile phone during and outside of class. While overall learning levels were quite low – students were primarily able to recognize words and solve simple math problems – they were sufficient to operate the mobile phone.

Using data collected two years after the end of the experiment in Niger (Aker et al 2012), this paper assesses whether the short-term learning gains observed in the ABC program persist several years after the program, as well as translate into other improvements in socio-economic outcomes. We find that the short-term reading gains from the ABC program persist several years' later: Students in ABC villages were 8 percentage points more likely to be able to read words and had reading z-scores that were .15 s.d. higher than those in non-ABC villages, with similar effects for men and women. Average math scores in ABC villages were the same as those in the standard adult education villages after the program, yet with stronger effects for female and younger students. The additional mobile phone instruction was also associated with improvements

in other socio-economic outcomes: ABC households engaged in more diverse income-generating strategies, were more likely to own a durable asset, had improved food security, were more likely to produce and sell a marginal cash crop and were more likely to save. Yet we are unable to disentangle whether these medium-term impacts on socio-economic outcomes are primarily due to improvements in learning, improved access to information by better manipulation of mobile phones, or some combination of the two.

Overall, this paper contributes to the existing literature on the medium-term effects of education programs. While there are a number of studies in this area, many of these are for primary education programs, thereby measuring skills acquisition and labor market outcomes in adulthood (Jacob et al 2010, Duflo 2003, Evans et al 2014). There is a smaller subset of papers that assess the impact of adult education programs on learning (Okech et al 2001, Royer et al 2004, Ortega et al 2008, Banerji et al 2016, Deshpande et al 2017) or welfare (Blunch and Portner 2011). Yet to our knowledge, only one of these studies focuses on adult education programs with a technology component (Deshpande et al 2017), and none of them assess whether short-term learning gains persist after the program.

The rest of this paper proceeds as follows. Section II describes the research design, whereas Section III outlines the data and estimation strategy. Section IV presents the core results, including the results on well-being and mechanisms, whereas Section V presents the alternative explanations. Section VI concludes.

## **II. Research Design**

The original experiment used in this paper is described in detail in Aker et al (2012), and is summarized here.<sup>1</sup> Starting in February 2009, an international non-governmental organization, Catholic Relief Services, implemented an adult education program in two regions of Niger. The program provided eight months of literacy and numeracy instruction to adults in 134 villages. The course cycle was for two years, and classes were held between February and June of each year, with a break between June and January due to the agricultural planting season. Classes were taught in the indigenous language of the village and were held five days per week for three hours per day. Each village had two literacy classes (separated by gender), with twenty-five students per class. Each class was taught by a community literacy teacher, with two teachers per village (Aker et al 2012).

The mobile phone intervention (ABC) was a module added to the basic adult education program. Students in ABC and non-ABC villages followed the same curriculum, but students in ABC villages also learned how to use a simple mobile phone (Aker et al 2012). In addition, a mobile phone was provided to groups of literacy participants in ABC villages, with one mobile phone per group of five people. The mobile phone module was introduced three months after the start of the adult education program. Teachers in ABC villages were instructed to teach the mobile phone module for approximately one hour during the weekly review class (Aker et al 2012)

The primary sample for the original experiment was 113 villages in two regions of Niger. The original experiment stratified villages by geographic divisions, and then randomly assigned villages to a cohort (to start adult education classes in 2009 or 2010)

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<sup>1</sup> The text in this section is paraphrased heavily from the original research design section of the Aker et al (2012) paper.

and either the basic (non-ABC) or the basic plus mobile-phone intervention (ABC). Those in the 2009 cohort ended their course in 2010, and those in the 2010 cohort ended in 2011.

Within each village, eligible students were identified for both cohorts during the baseline. Individuals had to be illiterate, willing to participate in the program and a member of either a formal or informal village-level producers' association. If there were more than fifty eligible applicants in a village, students were randomly chosen from among all eligible applicants in a public lottery.

As there was no pure control group in the original experiment, this paper estimates the causal effect of the ABC intervention as compared to the standard adult education intervention. As a result, we cannot estimate the causal impact of the adult education program, and therefore cannot answer questions pertaining to the returns to adult education.

### **III. Data and Estimation Strategy**

#### **A. Data**

The data we use in this paper come from two primary sources. First, we administered math and reading tests and use these to measure the impact of the program on learning outcomes several years after the program. Second, we conducted surveys on student and household characteristics. Before presenting our estimation strategy, we discuss each of these data sources in detail.

In 2013, two years after the end of the original program, we returned to one region (Dosso) to collect test score and household data. The focus on one region was purposeful: While there was not a statistically significant difference of the impact of the

ABC program by region in the original paper, the coefficients were larger in magnitude when focusing on the Dosso region (Aker et al 2012). As a result, our data collection efforts in 2013 focused solely on this region. Using the original list of all students in each village from 2009, we confirmed students' presence in the village in 2013. Among those present, we stratified by gender and chose a random sample of 16 students per village, 8 men and 8 women.

In our previous work, we assessed learning using the Ministry of Non-Formal Education's battery of tests. For this follow-up study, we used USAID's Early Grade Reading Assessment (EGRA) and Early Grade Math Assessment (EGMA) tests. EGRA is a series of timed tests that measure basic foundational skills for literacy acquisition: recognizing letters, reading simple words and phrases and reading comprehension (Dubeck and Gove 2015). Each task ranges from 60-180 seconds; if the person misses four answers in a row, the exercise is stopped. EGMA measures basic foundational skills for math acquisition: number recognition, comparing quantities, word problems, addition, subtraction, multiplication and division (Reubens 2009).

The EGRA and EGMA tests were our preferred survey instruments, as compared with the Ministry's untimed tests, for two reasons. First, most adult education programs are criticized for high rates of skills' depreciation. Yet skills' depreciation may be in part due to the low reading levels achieved by the end of traditional adult education programs, which cannot be easily captured in untimed tests. For example, the short-term memory required to store deciphered material is brief, lasting 12 seconds and storing 7 items (Abadzi 2003). Thus, "Neoliterates must read a word in about 1-1.5 second (45-60 words

per minute) in order to understand a sentence within 12 seconds (Abadzi 2003).<sup>2</sup> Thus, the EGRA timed tests allow us to determine whether participants in adult education classes are attaining the threshold required for sustained literacy acquisition. Second, the tests offer a great deal of precision in terms of measuring the skills that contribute to reading and math acquisition, capturing more nuanced levels of variation in learning (Dubeck and Gove 2015). While this makes it difficult for us to directly compare learning levels between 2011 and 2013, the EGRA tests can be easily converted into the Ministry's reading levels. This mapping is more challenging for the math tests.

In addition to the reading and math tests, students participated in a household survey, which asked questions about socio-demographic characteristics, assets, agricultural production and marketing, livestock production and food security. These data are used to measure the longer-term impacts of the program on different indicators household assets, agricultural and savings behavior, and food security status.

Attrition is a concern in most studies, especially those that return after the end of the program and amongst populations who engage in seasonal migration. Table 2A formally tests whether there is differential attrition by ABC status for the follow-up survey rounds in 2013. The rate of attrition was 5 and 4 percent in non-ABC and ABC villages, respectively, without a statistically significant difference between the two (Table 2A). The primary reasons for attrition in the 2013 follow-up sample were migration, illness and death. Of these, the mortality rates were higher in non-ABC villages as compared with ABC villages (conditional on attrition), with a statistically significant

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<sup>2</sup>This speed corresponds to oral-reading U.S. norms for first grade children. However, this is often not attained in literacy classes. For example, studies in Burkina Faso indicate that most literacy graduates need 2.2 seconds to read a word and are correct only 80-87 percent of the time (Abadzi 2003).

difference between the two. Yet the magnitude of the mortality rate was small, representing less than 1.5 percent of the overall sample. Looking at the determinants of attrition (Table 2B), ABC status, gender and residence were not individually or jointly significant determinants of attrition (Table 2B), suggesting that attrition is not a primary concern in our context.

### **B. Balance of Time-Invariant Characteristics**

The original experiment yielded groups that were similar along observable characteristics prior to the program (Aker et al 2012). Table 1A shows this balance from the original experiment for the Dosso region, the focus of this paper. Since the program occurred between 2009 and 2011, we would expect time-varying characteristics to change as a result of the program. However, we can test for whether the original experimental balance remains among our new sample by testing for the equality of means for time-invariant characteristics between ABC and non-ABC villages. Table 1B presents these results. Overall, there are no statistically significant differences in time-invariant characteristics between the ABC and non-ABC villages.

### **C. Estimation Strategy**

In order to measure the medium-term impacts of the ABC program, we use the following estimation strategy:

$$(1) \quad test_{icv} = \alpha + \beta_1 ABC_v + \gamma female_i + \delta cohort_v + \theta_R + \epsilon_{iv}$$

Where  $test_{icv}$  is the Early Grade Reading Assessment (EGRA) or Early Grade Math Assessment (EGMA) score attained by student  $i$  two years after the program,  $ABC_v$  is an indicator variable for whether individual  $i$  was assigned to the ABC or non-ABC intervention between 2009 and 2011,  $\theta_R$  are fixed effects that indicate the randomization

strata, *female* is a binary variable for the student's gender, and  $cohort_v$  is the year the village started in the adult education program (equal to 1 if 2010 cohort), since the program was phased in over time. We cluster standard errors at the village level and correct for heteroscedasticity.

A key assumption in the identification of the medium-term treatment effect is that the independence assumption holds. This assumption appears to be valid, as presented in the Table 1. Additional threats to the validity of the treatment effect are discussed in Section V.

#### **IV. Results**

##### **A. Medium-Term Effects of the Program**

Table 3A presents the results of equation (1) on reading scores two years' after the end of the program. Overall, the results suggest that the short-term learning gains persisted: ABC students were 8 percentage points more likely to be able to read any word and were able to correctly read a word as compared to their non-ABC counterparts, with a statistically significant effect at the 5 percent level (Panel A). These results are also robust to the use of reading z-scores: ABC students achieved reading levels that were .15 s.d. higher as compared to non-ABC students, with a statistically significant effect at the 5 percent level (Column 3). While ABC students reading automaticity and comprehension were also higher as compared to those in non-ABC villages, these effects are not statistically significant at conventional levels (Columns 4 and 5). Although the EGRA tests do not map perfectly to the previous Ministry tests, the results are robust to this mapping: students in ABC villages read at a level that was almost half a point higher than those in non-ABC villages two years after the program, a 25% increase. Similar to

the short-term effects of the program, there are no differential effects by gender or age (Panels B and C).<sup>3</sup>

Table 3B shows the results of the estimation of equation (1) for math scores. While average math scores – including number identification, simple addition and subtraction and a composite math score - are higher in ABC villages, none of these differences are statistically significant at conventional levels (Panel A). Yet the heterogeneous results by gender and age suggest that the ABC program enabled women and younger students to continue to correctly identify numbers several years after the program (Panels B and C, Columns 1 and 2). In particular, for those students younger than 39 years of age, the ABC program helped them to retain their number identification and general math skills (Panel C, Columns 3-5).

In our previous work, the ABC program increased average math scores by .25 s.d. relative to the standard adult education program in the short-term, with no differential effects by gender or age. What could potentially explain the absence of an average effect several years after the program? First, ABC students' math skills could have depreciated during this time, thus returning to non-ABC levels at the end of the program. This is possible, as, other than identifying numbers, mobile phone usage – which was the crux of the ABC program -- would not necessarily assist students in using and hence retaining more complicated mathematical operations, such as addition and subtraction. In addition, our previous work showed strong rates of depreciation in learning six months' after the

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<sup>3</sup>While not shown, there are also no differential effects of the ABC program by language (Hausa or Zarma); this is particularly relevant for adult education programs, as the language's transparency can affect how easy it is for adults to learn and retain reading skills. Finally, given the small number of clusters (35 in each treatment group), we corrected the standard errors using a Wild Cluster bootstrap, and found similar results.

program, and the ABC program was not effective in reducing the rate of depreciation (Aker et al 2012).<sup>4</sup>

An alternative explanation for the absence of average effects on math scores is the way in which the math tests were administered in 2013. In particular, the EGRA tests allowed for addition and subtraction tasks to be completed *verbally*. Hence, the 2013 tests did not capture ABC students' ability to decode written mathematical problems, as was the case in 2011 math tests from our previous work.

### **B. Skills Depreciation after the Program**

The results in Tables 3A and 3B suggest that the additional skills that ABC students initially attained became smaller over time, especially for math. This could be due to the typical depreciation observed in adult education programs (Abadzi 2003).<sup>5</sup> Yet simply comparing math and reading test score levels after the program ignores students' learning at the end of the program, which should be highly correlated with persistence.

A common method to account for persistence is to use the value-added model – either restricted or lagged – that controls for the previous period's test scores.<sup>6</sup> Andrabi et al (2011) show that a standard estimation of this model – especially when previous test scores are endogenous – yields biased estimates of both the persistence coefficient and

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<sup>4</sup>While non-ABC students could have also improved their math skills after the end of the program, thereby “catching up” to ABC students, this seems unlikely, as there were not additional or new adult education programs in these villages between 2011 and 2013.

<sup>5</sup>An interesting question is, however, whether the loss in learning is smaller or larger for ABC learners relative to what would be expected. This could be because students continue to use their skills more often or even continue expanding their skillset due to interactions with the phone.

<sup>6</sup>In 2011, we used Ministry tests, which mapped students' learning gains from 0 to 7, based upon a series of 6 tasks. In 2013, the EGRA and EGMA tests involved a series of timed tasks. While it is straightforward to map the 2013 reading test scores to the 2011 reading tests, this is less straightforward for math. Nevertheless, there are some tasks that are similar across both tests, which allows us to map the 2013 results to the 2011 levels.

the treatment effect. In order to correct for this bias, they suggest using panel data methods over three data collection rounds.

In this paper, we have data from two periods – the program endline data (in 2011) and the follow-up data (in 2013), and thus cannot use the panel data methods suggested by Andrabi et al (2011). While data are also available from the original baseline conducted in 2009, these data are not informative, as almost all baseline reading and math levels were zero. Thus, to test whether ABC slowed the loss in skills, we use the 2011 and 2013 data from the same sample of students in the following value-added specification:<sup>7</sup>

$$(2) \quad test_{icv2013} = \alpha + \beta_1 ABC_v + \beta_2 test_{icv2011} + X_i \gamma + \delta cohort_v + \theta_R + \varepsilon_{iv}$$

where  $\beta_1$  is the treatment effect and  $\beta_2$  is the persistence coefficient.  $\beta_1$  is the medium-term effect of the ABC program, accounting for the expected depreciation.

Since the 2011 test scores were affected positively by the ABC program, both student-specific heterogeneity in learning and measurement error will bias  $\beta_2$ , thereby yielding biased estimates in both the treatment and persistence coefficients. To partially address this bias, we use the control function approach, similar to that outlined by Petrin and Train (2002). At the first stage, we regress 2011 test scores on village-level fixed effects, which also captures the ABC treatment variable, in the following specification:

$$(3) \quad test_{icv2011} = \theta_V + u_{iv}$$

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<sup>7</sup>There are two primary value-added specifications in the education literature: the restricted value-added model and the lagged value-added model. The former assumes that “lagged achievement contributes cumulatively without loss to future achievement”, whereas the latter allows the contribution of previous achievement may decay (Andrabi et al 2011). A variety of specifications using the lagged value-added model suggest a depreciation coefficient of .6 (Andrabi et al 2011).

At the second stage, we include the estimated residual from the first stage as a covariate in the following specification:

$$(4) \quad test_{icv2013} = \alpha + \beta_1 ABC + \delta \hat{u}_{iv2011} + \varepsilon_{iv}^8$$

The above approach corrects for any potential bias in the treatment effect, but not for the persistence effect. While this is potentially problematic, Andrabi et al. (2011) find that the bias from heterogeneity is larger than the bias from measurement error. This may also be the case for the tests studied here, which are designed to measure key building blocks of literacy and of numeracy. Omitted heterogeneity leads to an overestimate in the persistence coefficient and an underestimate of the treatment effect. In our case, this would suggest that the impacts of the ABC program on skills retention are likely to be underestimated. Nevertheless, we recognize that this approach does not fully remove all potential selection bias induced from the value-added specification.

Table 4 shows the results of these specifications, first using the standard OLS specification in equation (2) (Panel A). Using standard methods, reading and math test scores – as measured by the Ministry of Non-Formal Education’s battery of tests -- are higher in ABC villages than non-ABC villages after controlling for prior achievement, although only one of these coefficients is statistically significant at conventional levels. Yet there is differential persistence: the coefficient on the lagged dependent variable ranges from .31 (for math) to .78 (for reading).

Panel B attempts to correct for potential bias in the treatment effect by using the control function approach outlined above. While the persistence coefficients remain

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<sup>8</sup>While the standard errors presented are clustered at the village level, they are robust to the approach used by Petrin and Train, which involves first estimating robust standard errors clustered at the village level, drawing bootstrap samples, identifying how much the estimated coefficient varies due to the estimation of village fixed effects in the first stage, and adding this to the estimated variance from the clustering.

fairly stable – ranging between .30 and .75, the coefficients on the *ABC* variable are strongly positive and statistically significant for reading, and marginally statistically significant for math. This supports the results provided in Table 3, suggesting that the impacts of the *ABC* program on learning outcomes persisted, especially for reading, two years after the program.

### **C. Mechanisms**

What are the mechanisms through which the short-term learning gains in learning persisted? And can these gains be attributed to learning or continued practice with the mobile phone? In Aker et al (2012), the authors found that the primary mechanism for short-term learning gains was increased mobile phone usage: students living in *ABC* villages were more likely to use the mobile phone in more active ways, both during classes and seven months after the end of the program. Table 5 suggests that this practice continued after the end of the program. While the *ABC* program was not associated with a higher likelihood of mobile phone usage or increased spending on mobile phone airtime, *ABC* students used mobile phones in more active ways two years after the program. They were significantly more likely to make calls, write and receive SMS and send and receive airtime as compared with their non-*ABC* counterparts, all of which require literacy and numeracy skills to different degrees. Yet this did not translate into changes in students' social or professional communications: *ABC* students were equally likely to talk with relatives or business contacts as their non-*ABC* counterparts. This suggests that *ABC* students may have increased their communication in less expensive ways, such as SMS. Overall, similar to the short-term results, Table 5 provides evidence

that ABC students continued to use mobile phones in more active ways after the program, allowing for sustained opportunities to learn.

#### **D. The Impact of the ABC Program on Household Socio-Economic Outcomes**

While the previous results suggest that the short-term learning gains of the ABC program persisted two years after the program, a key question for many adult education programs – especially those using technology -- is whether this translates into improvements in adults' standard of living. While Aker and Ksoll (2015) found that students in ABC villages were more likely to grow a diverse set of crops immediately after the program, there were no other observable improvements in household socio-economic outcomes. Yet, the returns to educational investments may take several years to manifest. As our original experiment did not have a pure control group, we are only able to assess the impact of the ABC program (as compared with the standard adult education program) on these indicators, rather than the returns to the adult education program alone. As such, we are unable to determine whether the impacts on socio-economic outcomes are due to increased literacy skills, the increased use of mobile phones or some combination of the two.

Table 6 shows the results of estimating equation (1) on a number of outcomes, including the household's sources of income and self-reported income, assets, food security, savings, agricultural practices and intra-household decision-making. As we do not have a full consumption and expenditure module, we are unable to assess the impacts of the ABC program on welfare, but use these indicators as a proxy. In addition, it is important to note that some of the food security measures – namely, the household diet

diversity score – were only asked of those respondents who had prepared the food in the past 24 hours, and therefore includes approximately one-third of the sample.

Overall, the results suggest that ABC students were faring better two years after the program, using a broad variety of indicators. ABC households were also less likely to experience food insecurity (Panel A): Conditional on having gone without food for a day in the past year, households were 13 percentage points less likely to go hungry frequently. In terms of dietary quality, ABC households had diet diversity that was 10% higher as compared to non-ABC villages, and were 11 percentage points more likely to consume 6 food groups (out of 12) over the past 24 hours, a key threshold for improved nutritional status. ABC households also owned more assets than non-ABC households, primarily bikes (Panel A).

These changes in food security and asset ownership are potentially explained by households' changes in their income-generating strategies, agricultural activities and savings behavior. Households in ABC villages were using .15 more strategies to earn income, primarily trade and small business, although they did not report any differences in self-reported income (Panel B). In terms of agriculture (Panel C), ABC households were 7 percentage points more likely to apply fertilizer, and produced and sold more types of crops, primarily *gombo*, a traditional cash crop grown by women. This latter finding is consistent with the short-term results of the impact of the ABC program on agricultural outcomes (Aker and Ksoll 2016).<sup>9</sup>

In the medium-term, the ABC program was also associated with impacts on households' savings behavior: households in ABC villages were 11 percentage points

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<sup>9</sup>As we do not have detailed data on gombo production, the amount sold or sales prices, we are unable to determine whether this led to increases in income earned from the sale of gombo.

more likely to belong to a savings group, and felt less obligated to share money with friends or family members, a proxy for some control over finances (Panel D).

A key interest of many adult education programs is improving women's empowerment (Banerji et al 2016). Panel E assesses the impact of the ABC program on intra-household decision-making, using a composite indicator measuring decision-making on children's schooling, crop choice and health decisions. The responses are coded so that it equals "one" if the respondent decides alone, zero if jointly. Thus, a reduction in the score would imply that the household was making decisions jointly. Restricting the sample to only married couples, the ABC program did not have a strong or statistically significant impact on intra-household decision-making.

A key question underlying these results is whether the observed results are due to improved *learning* or increased usage of the mobile phone. While we do not have the data to disentangle these effects, the results on mobile phone ownership and usage suggest that ABC households were not using the mobile phone for agricultural trade or to receive remittances from migrants. Thus, this provides suggestive evidence that some of the observed impacts can be attributed to improved learning via mobile phone instruction, and not only via increased mobile phone usage.

## **V. Alternative Explanations**

There are several potential threats to the validity of the above findings. While there was no differential attrition by ABC status several years' after the program, it is possible that the ABC program could have changed households' migration patterns, either for the student or for other household members. If ABC household members were more likely to engage in seasonal off-migration, thereby contributing to household income, this could

potentially explain the observed impacts on assets and food security. While this would not invalidate the observed impacts – as these results could be attributed to the ABC program – it would change the interpretation of our findings. Yet our results suggest that, two years after the program, neither the extensive nor intensive margin of migration differed by ABC status (not shown). Thus, this does not appear to be driving the results.<sup>10</sup>

A second threat to the validity of these findings is spillovers. For example, if students in ABC villages directly or indirectly affected the learning or socio-economic outcomes of households in non-ABC villages, then this could over or underestimate our effects. While we have no evidence that spillovers occurred, we hypothesize that such spillovers would result in a lower bound of the impact on learning outcomes, although we are unable to sign the potential impact of spillovers on other socio-economic outcomes.

Third, while we do not find evidence of differential attrition in the initial census exercise, and we may be concerned that any differences in attrition may affect our findings. However, the rate of attrition (non-response) in the actual selected sample was extremely low, less than 1%. Estimating Lee bounds correcting for non-differential attrition in the selected sample shows that all of the lower bounds remain positive and statistically significant for reading, and positive for math (not shown). As expected, all of the upper bounds are positive and statistically significant.

## **VI. Conclusion**

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<sup>10</sup>There is no differential attrition between the ABC and non-ABC villages, and attriters do not differ along observable dimensions. Thus, we do not feel that differential attrition could be driving the results.

This paper reports the impacts of a mobile phone-enhanced adult education program in Niger two years after the program. Overall, we find that short-term learning gains persisted two years after the program, as students in ABC villages had higher test scores than their non-ABC counterparts, primarily for reading, with heterogeneous effects for math. These effects are primarily explained by more active mobile phone usage after the end of the program, which enabled households to continue practicing their reading skills via texting and sending airtime.

The ABC program also resulted in improvements in households' other socio-economic outcomes, namely food security and asset ownership. These improvements seem to be primarily correlated with changes in income generating strategies, agricultural activities and savings behavior: ABC households were more likely to engage in trade, produce and sell more marginal cash crops and engage in savings groups. While we are unable to disentangle whether these improvements are primarily attributed to the improved learning or increased mobile phone usage, they are not due to increased mobile phone ownership, which did not differ between the two groups. In the future, a possible way to disentangle the mechanisms for the observed welfare impacts would be to implement a cross-cutting experiment in which study participants are assigned to one of four groups: a control group, an adult education program, a mobile phone program (without an adult education component) and the ABC program. Assuming that the mobile phone treatment would not significantly increase educational outcomes in the absence of an adult education program, then any welfare impacts observed in this group would be due to the mobile phone. Nevertheless, in a mobile phone cash transfer program in

Niger, Aker et al (2016) found that simply providing a mobile phone (with cash) did not lead to any additional welfare improvements as compared with the cash only treatment.

While many education programs are often plagued by high rates of skills depreciation after the program, these results suggest that simple information technology may be able to temper such depreciation, especially if it can be used in daily life and reinforce the skills learned in class. Nevertheless, such technology must be readily accessible to students after the program, and something that they can use themselves – which is not always the case with information technology that is classroom-based.

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**Table 1A: Baseline Means Comparison for Dosso**

	(1)	(2)	(3)	(4)
	<b>Non-ABC Mean (s.d.)</b>	<b>ABC Mean (s.d.)</b>	<b>Conditional Difference Coeff (s.e.) Coeff (s.e.)</b>	<b>Number of observations</b>
Age of respondent	40.50 (13.59)	40.30 (11.45)	-0.24 (1.29)	532
Respondent is household head (1=Yes, 0=No)	0.57 (0.50)	0.57 (0.50)	0.00 (0.03)	530
Respondent is married (1=Married, 0=Single, divorced, widowed)	0.93 (0.25)	0.93 (0.26)	-0.01 (0.02)	532
Member of Hausa ethnic group	0.79 (0.41)	0.71 (0.45)	-0.06 (0.05)	532
Number of household members	9.06 (4.44)	9.19 (4.48)	0.22 (0.35)	522
Percentage of children (less than 15) with some education	0.33 (0.28)	0.32 (0.28)	-0.01 (0.03)	532
Number of asset categories owned	5.04 (1.77)	5.20 (1.81)	0.10 (0.14)	532
Household experienced drought in the past year	0.56 (0.50)	0.55 (0.50)	0.00 (0.04)	529
Household owns mobile phone (1=Yes, 0=No)	0.38 (0.49)	0.40 (0.49)	0.00 (0.04)	494
Respondent has used mobile phone since last harvest (1=Yes, 0=No)	0.65 (0.48)	0.68 (0.47)	0.02 (0.05)	531
Respondent has used mobile phone to make calls	0.71 (0.45)	0.74 (0.44)	0.04 (0.05)	355
Respondent has used mobile phone to receive calls	0.91 (0.29)	0.90 (0.29)	-0.01 (0.03)	355

Notes: Column 1 presents the unconditional mean for non-ABC villages, Column 2 presents the unconditional mean for ABC villages. Column 3 reports the coefficient from a regression of the dependent variable on an indicator variable for ABC, cohort and sub-region fixed effects to account for randomization. Thus, Column (3) is not exactly equal to the difference between Columns 1 and 2. The sample is a random sample (12 households) from among all program participants in each village in the Dosso region. Huber-White standard errors clustered at the village level presented in parentheses. \*\*\*, \*\*, \* denote statistical significance at the 1, 5, 10 percent levels, respectively.

**Table 1B. Means Comparison of Socio-Demographic Characteristics**

	(1)	(2)	(3)
	<b>Non-ABC</b>	<b>ABC</b>	<b>Difference in Means (Conditional)</b>
	Mean (s.d.)	Mean (s.d.)	Coeff (s.e.)
Age of respondent	41.67 (13.66)	41.12 (12.92)	-0.45 (1.25)
Born in Dosso	0.97 (0.18)	0.98 (0.15)	0.01 (0.01)
Respondent is married (1=Married, 0=Single, divorced, widowed)	0.93 (0.26)	0.91 (0.29)	-0.03 (0.02)
Speak Hausa	0.90 (0.30)	0.88 (0.32)	-0.03 (0.02)
Number of household members	11.03 (4.73)	10.51 (4.37)	-0.45 (0.33)
Mean age of household members	19.41 (5.60)	19.77 (5.03)	0.24 (0.38)
Mean gender of household members	0.50 (0.14)	0.50 (0.14)	0.00 (0.01)

Notes: Column 1 presents the unconditional mean for non-ABC villages, Column 2 presents the unconditional mean for ABC villages, and Column 3 reports the coefficient from a regression of the dependent variable on an indicator variable for ABC, cohort and sub-region fixed effects, which account for randomization. Thus, Column (3) is not exactly equal to the difference between Columns 1 and 2. The sample is a random sample in Table A1, in that it is a random sample of 16 households per village from among all program participants in each village region. Huber-White standard errors clustered at the village level presented in parentheses. \*\*\*, \*\*, \* denote statistical significance at the 1, 5, and 10 percent levels, respectively.

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(4)

**Number of  
Observations**

1,115

1,122

1,122

1,121

1,122

1,122

1,122

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<b>Table 2A. Attrition</b>				
	(1)	(2)	(3)	(3)
	<b>Non-ABC</b>	<b>ABC</b>	<b>Difference in Means (Conditional)</b>	<b>Number of Observations</b>
<i>Panel A. Attrition</i>	Mean (s.d.)	Mean (s.d.)	Coeff (s.e.)	
Attrition	0.05	0.04	-0.00	3,461
	(0.21)	(0.21)	(0.01)	
<i>Panel B. Causes of Attrition</i>				
Illness	0.04	0.01	-0.04	157
	(0.19)	(0.11)	(0.04)	
Migration	0.42	0.45	0.05	157
	(0.50)	(0.50)	(0.09)	
Death	0.22	0.12	-0.09*	157
	(0.41)	(0.32)	(0.05)	
Other (visiting friends, baptism)	0.33	0.42	0.08	157
	(0.47)	(0.50)	(0.10)	

Notes: Column 1 presents the unconditional mean for non-ABC villages, Column 2 presents the unconditional mean for ABC villages. Column 3 reports the coefficient from a regression of the dependent variable on an indicator variable for ABC, cohort and sub-region fixed effects to account for randomization. Thus, Column (3) is not exactly equal to the difference between Columns 1 and 2. Results are robust to omitting the sub-region fixed effects. Huber-White standard errors clustered at the village level presented in parentheses. \*\*\*, \*\*, \* denote statistical significance at the 1, 5, 10 percent levels, respectively.

**Table 2B. Determinants of Attrition**

	Coeff (s.e.)	Coeff (s.e.)
ABC	-0.00	-0.00
	(0.01)	(0.01)
Female		-0.02**
		(0.01)
Doutchi		-0.01
		(0.02)
Number of Observations	3,461	3,461
R-squared	0	0.01
P-value of joint F-test		0.22

Notes: Column 1 estimates a regression of a binary variable for attrition on an ABC treatment indicator. Column 2 estimates a regression of a binary variable for attrition on ABC treatment, cohort and stratification fixed effects, female and Doutchi. The sample consists of all of the program participants (50 per village) from the original ABC program in each village in the Dosso region. Huber-White standard errors clustered at the village level presented in parentheses. \*\*\*, \*\*, \* denote statistical significance at the 1, 5, 10 percent levels, respectively.

**Table 3. Impact of ABC on Reading Scores**

	(1)	(2)	(3)	(4)	(5)
	Any Word Read	Number of Words Read	Z-Score Number of Words Read	Number of Words Read/Second	Reading Comprehension
<i>Panel A: All Students</i>					
ABC Village	0.08** (0.03)	2.43** (1.18)	0.15** (0.07)	0.12 (0.08)	0.13 (0.09)
Female	Yes	Yes	Yes	Yes	Yes
2010 Cohort	Yes	Yes	Yes	Yes	Yes
Strata Fixed Effects	Yes	Yes	Yes	Yes	Yes
Mean of Non-ABC Villages	0.26	7.28	0	0	0.77
Number of observations	1,124	1,124	1,124	1,124	1,124
R-squared	0.12	0.12	0.11	0.12	0.12
<i>Panel B: Gender</i>					
ABC Village	0.07 (0.04)	2.28 (1.91)	0.14 (0.12)	0.10 (0.12)	0.07 (0.16)
ABC*Female	0.03 (0.06)	0.28 (2.18)	0.02 (0.14)	0.05 (0.14)	0.09 (0.18)
Female	-0.24*** (0.04)	-9.24*** (1.30)	-0.58*** (0.08)	-0.53*** (0.08)	-0.91*** (0.12)
2010 Cohort	Yes	Yes	Yes	Yes	Yes
Strata Fixed Effects	Yes	Yes	Yes	Yes	Yes
Number of observations	1,124	1,124	1,124	1,124	1,124
R-squared	0.12	0.11	0.12	0.11	0.11
<i>Panel C: Age</i>					
ABC Village	0.04 (0.04)	0.74 (1.41)	0.05 (0.09)	0.00 (0.02)	-0.04 (0.13)
ABC*Young	0.07 (0.05)	2.18 (1.81)	0.14 (0.11)	0.02 (0.02)	0.17 (0.17)
Young (<=39 years)	0.15*** (0.04)	4.62*** (1.12)	0.29*** (0.07)	0.04*** (0.01)	0.34*** (0.12)
Female	Yes	Yes	Yes	Yes	Yes
2010 Cohort	Yes	Yes	Yes	Yes	Yes
Strata Fixed Effects	Yes	Yes	Yes	Yes	Yes
Number of observations	1,034	1,034	1,034	1,034	1,034
R-squared	0.12	0.12	0.12	0.12	0.11

Notes: This table estimates the impact of the ABC program on a variety of reading indicators. ABC is equal to 1 if a student lived in a ABC village between 2009 and 2011, 0 otherwise. All regressions control for cohort and randomization. Huber-White standard errors clustered at the village level are in parentheses.

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(6)
Original Reading Level (0-7)
0.43**
(0.19)
Yes
Yes
Yes
1.63
1,124
0.12
0.39
(0.28)
0.21
(0.36)
-1.48***
(0.24)
Yes
Yes
1,124
0.12
0.22
(0.24)
0.39
(0.29)
0.83***
(0.20)
Yes
Yes
Yes
1,034
0.13

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fixed effects.

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**Table 3B. Impact of ABC on Math Scores**

	(1)	(2)	(3)	(4)	(5)	(6)
	Numbers correctly identified (out of 20)	Z-Score Number Identification	Oral Addition and Subtraction	Z-Score of Addition and Subtraction	Math Composite Z- Score	Original Math Level (0-7)
<i>Panel A: All Students</i>						
ABC Village	0.79 (0.53)	0.10 (0.07)	0.17 (0.13)	0.08 (0.06)	0.10 (0.07)	0.26 (0.16)
Female	Yes	Yes	Yes	Yes	Yes	Yes
2010 Cohort	Yes	Yes	Yes	Yes	Yes	Yes
Strata Fixed Effects	Yes	Yes	Yes	Yes	Yes	Yes
Mean of Non-ABC Villages	11.45	0	7.7	0	0	2.59
Number of observations	1,124	1,124	1,124	1,124	1,124	1,124
R-squared	0.29	0.12	0.14	0.32	0.28	0.14
<i>Panel B: Gender</i>						
ABC Village	0.02 (0.58)	0.00 (0.07)	-0.00 (0.16)	-0.00 (0.07)	0.01 (0.07)	0.12 (0.21)
ABC*Female	1.56* (0.87)	0.20* (0.11)	0.34 (0.30)	0.16 (0.15)	0.19* (0.11)	0.28 (0.25)
Female	-8.57*** (0.56)	-1.08*** (0.07)	-1.15*** (0.21)	-0.55*** (0.10)	-1.06*** (0.08)	-1.36*** (0.16)
2010 Cohort	Yes	Yes	Yes	Yes	Yes	Yes
Strata Fixed Effects	Yes	Yes	Yes	Yes	Yes	Yes
Number of observations	1,124	1,124	1,124	1,124	1,124	1,124
R-squared	0.29	0.29	0.12	0.12	0.29	0.09
<i>Panel C: Age</i>						
ABC Village	-0.57 (0.68)	-0.07 (0.09)	-0.07 (0.20)	-0.03 (0.10)	-0.06 (0.09)	-0.02 (0.21)
ABC*Young	1.87** (0.83)	0.24** (0.10)	0.40* (0.22)	0.19* (0.11)	0.22** (0.11)	0.49* (0.26)
Young (<=35 years)	2.40*** (0.55)	0.30*** (0.07)	0.22 (0.16)	0.11 (0.08)	0.31*** (0.07)	0.15 (0.17)
Female	Yes	Yes	Yes	Yes	Yes	Yes
2010 Cohort	Yes	Yes	Yes	Yes	Yes	Yes
Strata Fixed Effects	Yes	Yes	Yes	Yes	Yes	Yes
Number of observations	1,035	1,035	1,035	1,035	1,035	1,035
R-squared	0.34	0.34	0.12	0.12	0.33	0.10

Notes: This table estimates the impact of the ABC program on a variety of math indicators. ABC is equal to 1 if the student lived in a ABC village between 2009 and 2011, 0 otherwise. All regressions control for cohort and randomization fixed effects. Huber-White standard errors clustered at the village level are in parentheses.

**Table 4. Skills Depreciation**

	(1)	(2)	(3)
	Original Reading Level	Any Word Read	Original Math Level
<i>Panel A: OLS</i>			
ABC	0.25 (0.15)	0.08** (0.03)	0.18 (0.16)
Lagged Score (2011)	0.78*** (0.04)	0.46*** (0.04)	0.31*** (0.05)
Female	-0.58*** (0.15)	-0.14*** (0.03)	-0.91*** (0.15)
Cohort 2010	-0.07 (0.14)	-0.06* (0.03)	-0.10 (0.16)
Strata Fixed Effects	Yes	Yes	Yes
Number of observations	929	929	930
R-squared	0.43	0.35	0.14
<i>Panel B: Control Function Approach</i>			
ABC	0.59*** (0.22)	0.10*** (0.04)	0.27* (0.16)
Residual of Lagged Score (2011)	0.75*** (0.05)	0.44*** (0.04)	0.30*** (0.05)
Female	-0.63*** (0.15)	-0.15*** (0.03)	-0.92*** (0.15)
Cohort 2010	-0.41* (0.23)	-0.08** (0.04)	-0.13 (0.16)
Strata Fixed Effects	Yes	Yes	Yes
Number of observations	930	930	929

Notes: This table estimates the impact of the ABC program on a variety of reading and math indicators. ABC is equal to 1 if the student lived in a ABC village between 2009 and 2011, 0 otherwise. Panel A uses OLS, whereas Panel B includes the residuals from the first stage regression of 2011 test scores on village-level fixed effects as an additional control. All regressions control for cohort and randomization fixed effects. Huber-White standard errors clustered at the village level are in parentheses.

**Table 5. Mobile Phone Usage**

	(1)	(2)	(3)
	Mean (s.d.)	Coeff (s.e.)	Number of observations
Used mobile phone since last harvest	0.9 (0.31)	0.02 (0.02)	1119
Amount spent on mobile phone in past two weeks (CFA)	668 (1609)	6.39 (96.45)	1111
Made call	0.75 (0.43)	0.08*** (0.03)	756
Received call	0.96 (0.20)	0.01 (0.01)	756
Wrote SMS	0.06 (0.24)	0.06** (0.03)	756
Received SMS	0.24 (0.43)	0.10*** (0.03)	756
Transferred airtime	0.23 (0.42)	0.08*** (0.03)	756
Received airtime	0.37 (0.48)	0.09*** (0.03)	756
Talked with relative	0.93 (0.26)	0.03 (0.02)	754
Talked with commercial contact	0.09 (0.29)	0.02 (0.03)	754

Notes: This table estimates the impact of the ABC program on a variety of outcomes related to mobile phone ownership and usage. ABC is equal to 1 if the student lived in a ABC village between 2009 and 2011, 0 otherwise. All regressions control for cohort and randomization fixed effects. Huber-White standard errors clustered at the village level are in parentheses.

**Table 6. Household-Level Welfare Outcomes**

	(1)	(2)
	Mean (s.d.)	Coeff (s.e.)
<i>Panel A: Food Security and Asset Ownership</i>		
Household went a day without food in past year (0/1)	0.28 (0.45)	-0.02 (0.04)
past year	0.76 (0.43)	-0.13*** (0.04)
Household diet diversity (out of 12)	4.46 (2.00)	0.46* (0.27)
Reached household diet diversity threshold of 6	0.26 (0.44)	0.11* (0.06)
Z-score of number of household assets categories owned per capita	0.00 (1.00)	0.13 (0.08)
Bike ownership (0/1)	0.12 (0.33)	0.04* (0.02)
Mobile phone ownership (0/1)	0.87 (0.34)	0.03 (0.02)
<i>Panel B: Income Sources</i>		
Number of income sources (out of 12)	3.91 (1.26)	0.15** (0.07)
Self-reported income (of a scale of 5)	2.21 (0.86)	0.02 (0.06)
Household engaged in trade (0/1)	0.69 (0.46)	0.06** (0.03)
Household had a migrant	0.66 (0.47)	-0.04 (0.03)
<i>Panel C: Agriculture and Agricultural Practices</i>		
Number of crops produced	6.35 (1.53)	0.17 (0.14)
Number of crops sold	3.21 (0.18)	0.32* (0.18)
Gombo produced	0.63 (0.48)	0.07** (0.03)
Vegetables produced	0.29 (0.45)	0.07 (0.06)
Gombo sold	0.17 (0.38)	0.11** (0.04)
Vegetables sold	0.63	0.05

	(0.48)	(0.08)
Fertilized used	0.51	0.07*
	(0.50)	(0.04)
Improved seeds used	0.39	0.01
	(0.50)	(0.03)
Irrigation used	0.19	0.05
	(0.40)	(0.04)

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*Panel D: Savings and Risk Preferences*

Belonged to a savings group	0.35	0.11***
	(0.48)	(0.03)
Number of different savings mechanisms used	3.03	0.08
	(1.14)	(0.06)
Would like to spend less	0.45	0.08***
	(0.50)	(0.03)
Obligated to share money with a friend or family member	0.91	-0.04**
	(0.30)	(0.02)
Impatient	0.81	-0.02
	(0.39)	(0.02)
Risk averse	0.68	0.01
	(0.47)	(0.02)

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*Panel E: Intra-Household Decision-Making*

Z-score of index of autocratic decision-making	0.00	-0.04
	(1.00)	(0.04)

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Notes: This table estimates the impact of the ABC program on a variety of outcomes related to household well-being equal to 1 if the student lived in a ABC village between 2009 and 2011, 0 otherwise. All regressions control for community randomization fixed effects. Huber-White standard errors clustered at the village level are in parentheses.

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(3)

Number of  
observations

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1,097

289

399

399

1,122

1,122

1,122

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1,119

1,122

1,119

1,122

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1,121

1,121

1,121

1,121

720

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1,120

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1,122

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